

## Example

Suppose a farmer wants to know measure the production of a certain crop say, wheat when he has a document on the records of soil fertility and rainfall of that place and the average humidity of that place/location. As we can understand that these 3 variables had a direct influence on the production of wheat. We think on a linear f of 3 variables explaining the variable, yield (Prod.)

$$\text{Yield} = a + b_1 \times \text{soil fertility} + b_2 \times \text{rainfall} + b_3 \times \text{humid}$$

Where  $a, b_1, b_2, b_3$  are to be determined.

Remember, the relationship might be non-linear and the constants might not be multiplicative of  $\#$ . But if both happens (linear is constant and multi existed in between (yield and the vars.)

## Mathematical development

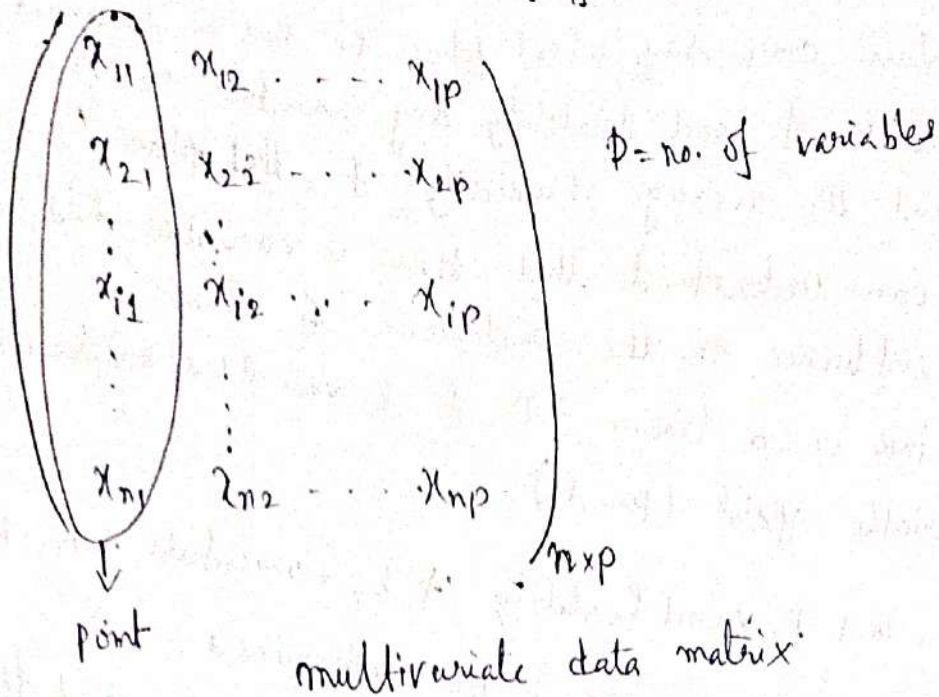
Suppose we have  $p$ -variables ( $x_1, x_2, x_3, \dots, x_p$ ) among which  $x_1$  being influenced by  $x_2, x_3, \dots, x_p$ . The relationship between  $x_1$  and  $x_2, x_3, \dots, x_p$  is assumed as linear as follows.

$$x_1 = a + b_2 x_2 + b_3 x_3 + \dots + b_p x_p \quad \text{--- (1)}$$

Clearly from the concept of bivariate regression,  $x_1$  is the dependent variable, while  $x_2, \dots, x_p$  ( $p-1$ ) are the independent variables and  $a, b_2, b_3, \dots, b_p$  are the constants to be determined.

(1) is called multiple regression line in a  $p$ -dimensional space.

we have a data-frame of size  $n \times p$ ,  
 where  $i$ th data point is  $(x_{i1}, x_{i2}, \dots, x_{ip})$ ;  $i = 1(1)n$



where  $\bar{x}_j = \text{mean for } j^{\text{th}} \text{ variable}$

$$\bar{x}_j = \frac{1}{n} \sum_{i=1}^n x_{ij}, \quad j = 1(1)p$$

$(s_{jj})$

$s_j^2 = \text{variance of } j^{\text{th}} \text{ variable}$

$$= \frac{1}{n} \sum_{i=1}^n (x_{ij} - \bar{x}_j)^2$$

$$s_{kj} = \text{cov}(x_k, x_j)$$

$$= \frac{1}{n} \sum_{i=1}^n (x_{ik} - \bar{x}_k)(x_{ij} - \bar{x}_j)$$

Our next objective is to estimate the coefficients.

Using Yule's dot notation we can write (1) in

$$x_1 = a_{1.23\dots p} + e_{1.23\dots p}$$

where  $x_1 = \text{observed data point for dependent variable}$ .

and  $x_{1,2,3 \dots p}$  = Predicted data point (from) of the dependent variable

Then,  $x_{1,2,3 \dots p} = a + b_2 x_2 + \dots + b_p x_p$

Moreover, the unexplained part of the dependent variable  $x_1$  is denoted by  $e_{1,2,3 \dots p}$  is error.

thus  $e_{1,2,3 \dots p} = x_1 - x_{1,2,3 \dots p}$

to estimate the constants, we use least square principle which minimizes the sum of squares of errors. (residual point)

$$\sum_{\alpha=1}^n e_{1,2,3 \dots p}^2 \quad \left| \quad e_{1,2,3 \dots p} = x_{\alpha 1} - a - b_2 x_{\alpha 2} - \dots - b_p x_{\alpha p} \right.$$

$$\Rightarrow L = \sum_{\alpha=1}^n (x_{\alpha 1} - a - b_2 x_{\alpha 2} - \dots - b_p x_{\alpha p})^2$$

The normal equations are,

$$\frac{\partial L}{\partial a} = 0 \Rightarrow -2 \sum_{\alpha=1}^n (x_{\alpha 1} - a - b_2 x_{\alpha 2} - \dots - b_p x_{\alpha p})$$

divide both side by n

$$\Rightarrow \bar{x}_1 - a - b_2 \bar{x}_2 - \dots - b_p \bar{x}_p = 0$$

$$\Rightarrow \boxed{a = \bar{x}_1 - b_2 \bar{x}_2 - \dots - b_p \bar{x}_p = 0}$$

$$\frac{\partial L}{\partial b_2} = 0 \quad \left| \quad \text{value input} \right.$$

$$\Rightarrow (-2) \sum_{\alpha=1}^n x_{\alpha 2} (x_{\alpha 1} - a - b_2 x_{\alpha 2} - \dots - b_p x_{\alpha p}) = 0$$

$$\Rightarrow \sum_{\alpha=1}^n x_{\alpha 2} x_{\alpha 1} - \bar{x}_1 \sum_{\alpha=1}^n x_{\alpha 2} + b_2 \bar{x}_2 \sum_{\alpha=1}^n x_{\alpha 2} + \dots + b_p \bar{x}_p \sum_{\alpha=1}^n x_{\alpha 2}$$

$$- b_2 \sum_{\alpha=1}^n x_{\alpha 2}^2 - \dots - b_p \sum_{\alpha=1}^n x_{\alpha 2} x_{\alpha p} = 0$$

Divide both side by n

$$\frac{1}{n} \sum_{\alpha=1}^n x_{\alpha 2} x_{\alpha 1} - \bar{x}_1 \frac{1}{n} \sum_{\alpha=1}^n x_{\alpha 2} + b_2 \bar{x}_2 \frac{1}{n} \sum_{\alpha=1}^n x_{\alpha 2} + \dots + b_p \bar{x}_p \frac{1}{n} \sum_{\alpha=1}^n x_{\alpha 2} - \frac{1}{n} \sum_{\alpha=1}^n x_{\alpha 2}^2 = 0$$

$$b_p \bar{x}_p \frac{1}{n} \sum_{\alpha=1}^n x_{\alpha 2} - \frac{b_2}{n} \sum_{\alpha=1}^n x_{\alpha 2}^2 - \dots - \frac{b_p}{n} \sum_{\alpha=1}^n x_{\alpha 2}^2 = 0$$

$$\Rightarrow S_{21} - \bar{x}_1 \cdot \bar{x}_2 + b_2 \bar{x}_2^2 + \dots + b_p \bar{x}_p \bar{x}_2 - b_2 S_{22} - \dots - b_p S_{2p} = 0$$

$$\Rightarrow S_{21} = b_2 S_{22} + \dots + b_p S_{2p} + \bar{x}_1 \bar{x}_2 - b_2 \bar{x}_2^2 - \dots - b_p \bar{x}_p \bar{x}_2$$

$$= b_2 S_{22} + \dots + b_p S_{2p} + \bar{x}_2 \cdot \hat{\alpha} \quad \text{--- (2)}$$

Linear in variable is sometimes confused with linear in parameters but as regression theory evolves from linear model, we always look into linearity in parameters or constants.

$$y = b_1 + b_2 x + \epsilon \rightarrow \begin{array}{l} \text{linear in constant} \\ \text{linear in variable} \end{array}$$

$$e^y = b_1 + b_2 x + \epsilon \rightarrow \begin{array}{l} \text{non-linear in variable} \\ \text{linear in parameter} \end{array}$$

$$y = b_1 + b_2 x^2 + \epsilon \rightarrow \begin{array}{l} \text{linear in constant} \\ \text{(polynomial regression)} \end{array}$$

$$y = \frac{1}{b_1} + e^{b_2 x} + \epsilon \rightarrow \begin{array}{l} \text{non-linear in constant} \\ \text{linear in variable} \end{array}$$

Any non-linear in constant model can be transformed into a linear regression format with suitable transformation in old constant.

$$r_{ij} = \frac{\sum_{\alpha=1}^n x_{\alpha i} x_{\alpha j}}{n \cdot s_i \cdot s_j}$$

$$r = \frac{\text{cov}(x_1, x_2)}{\sqrt{(x_1)} \sqrt{(x_2)}}$$

8.

$$s_{21} = \text{cov}(m_2, m_1)$$

$$r_{21} = \frac{\text{cov}(m_2, m_1)}{\sqrt{v(m_1) \cdot v(m_2)}}$$

$$\Rightarrow r_{21} \cdot s_{m_1} \cdot s_2 = s_{21}$$

(\*) Where,  $L = \sum_{\alpha=1}^n e_{1,2,3,\dots,p}^2$ ,  $e_{1,2,3,\dots,p} = \begin{matrix} x_{11} - a - b_2 x_{12} - \dots \\ \dots - b_p x_{1p} \\ \dots \\ x_{\alpha 1} - a - b_2 x_{\alpha 2} - \dots \\ \dots - b_p x_{\alpha p} \end{matrix}$

$$= \sum_{\alpha} (x_{\alpha 1} - a - b_2 x_{\alpha 2} - \dots - b_p x_{\alpha p})^2 \quad \left| \begin{matrix} e_{1,2,3,\dots,p} = x_{\alpha 1} - a - b_2 x_{\alpha 2} \\ \dots - b_p x_{\alpha p} \end{matrix} \right.$$

The normal equ<sup>n</sup> are,

$$\frac{\partial L}{\partial a} = 0 \Rightarrow (-2) \sum (x_{\alpha 1} - a - \dots - b_p x_{\alpha p}) = 0$$

$$\Rightarrow \bar{x}_1 - a - b_2 \bar{x}_2 - \dots - b_p \bar{x}_p = 0$$

$$\Rightarrow a = \bar{x}_1 - b_2 \bar{x}_2 - \dots - b_p \bar{x}_p$$

$$\frac{\partial L}{\partial b_2} = 0 \Rightarrow (-2) \sum x_{\alpha 2} \cdot (x_{\alpha 1} - a - \dots - b_p x_{\alpha p}) = 0$$

$$\Rightarrow \sum x_{\alpha 2} (x_{\alpha 1} - (\bar{x}_1 - b_2 \bar{x}_2 - \dots - b_p \bar{x}_p) - b_2 x_{\alpha 2} - \dots - b_p x_{\alpha p}) = 0$$

$$\Rightarrow \sum_{\alpha=1}^n x_{\alpha 2} x_{\alpha 1} - \bar{x}_1 \sum_{\alpha=1}^n x_{\alpha 2} + b_2 \bar{x}_2 \sum_{\alpha=1}^n x_{\alpha 2} + \dots + b_p \bar{x}_p \sum_{\alpha=1}^n x_{\alpha 2}$$

$$- b_2 \sum x_{\alpha 2}^2 - \dots - b_p \sum x_{\alpha p} x_{\alpha 2} = 0$$

$$\Rightarrow \frac{1}{n} \sum x_{\alpha 2} x_{\alpha 1} - \frac{\bar{x}_1}{n} \sum x_{\alpha 2} + b_2 \frac{\bar{x}_2}{n} \sum x_{\alpha 2} + \dots + b_p \frac{\bar{x}_p}{n} \sum x_{\alpha 2}$$

$$- \frac{b_2}{n} \sum x_{\alpha 2}^2 - \dots - \frac{b_p}{n} \sum x_{\alpha p} x_{\alpha 2} = 0$$

$$\Rightarrow \delta_{12} - b_2 \delta_{22} - b_3 \delta_{23} - \dots - b_p \delta_{2p} = 0$$

$$\Rightarrow \delta_{12} = b_2 \delta_{22} + b_3 \delta_{23} + \dots + b_p \delta_{2p}$$

similarly,

$$\delta_{13} = b_2 \delta_{23} + b_3 \delta_{33} + \dots + b_p \delta_{3p}$$

$$\vdots$$

$$\delta_{1p} = b_2 \delta_{2p} + b_3 \delta_{3p} + \dots + b_p \delta_{pp}$$

System of non-homogeneous linear equations

Remark: 'Linear' in regression means linear in parameters (constant).

► Any non-linear in constant model can be transformed into linear regression with suitable transformation in old constant.

► Linear regression is considered as fundamental model.

$$a = \bar{x}_1 - b_2 \bar{x}_2 - \dots - b_p \bar{x}_p$$

$$\delta_{12} = b_2 \delta_{22} + b_3 \delta_{23} + \dots + b_p \delta_{2p}$$

$$\delta_{13} = b_2 \delta_{23} + b_3 \delta_{33} + \dots + b_p \delta_{3p}$$

$$\vdots$$

$$\delta_{1p} = b_2 \delta_{2p} + b_3 \delta_{3p} + \dots + b_p \delta_{pp}$$

$$[\delta_{ij} = \text{Cov}(x_i, x_j)]$$

$$i = 1(1)p$$

$$\delta_{ii} = V(x_i)$$

$$\rightarrow \textcircled{2} = \delta_i^2$$

Define sample variance covariance matrix as,

$$S_{p \times p} = \begin{pmatrix} v(x_1) & \text{cov}(x_1, x_2) & \text{cov}(x_1, x_3) & \dots & \text{cov}(x_1, x_p) \\ \text{cov}(x_2, x_1) & v(x_2) & \text{cov}(x_2, x_3) & \dots & \text{cov}(x_2, x_p) \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ \vdots & \vdots & \vdots & \vdots & \vdots \\ \text{cov}(x_p, x_1) & \text{cov}(x_p, x_2) & \text{cov}(x_p, x_3) & \dots & v(x_p) \end{pmatrix}$$

$$= \begin{pmatrix} \delta_{11} & \delta_{12} & \delta_{13} & \dots & \delta_{1p} \\ \delta_{21} & \delta_{22} & \delta_{23} & \dots & \delta_{2p} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ \vdots & \vdots & \vdots & \vdots & \vdots \\ \delta_{ip} & \delta_{2p} & \delta_{3p} & \dots & \delta_{pp} \end{pmatrix}_{p \times p}$$

Note that:

$S$  is a symmetric matrix.

(2) can be rewritten in terms of system of non homogeneous equ<sup>n</sup>.

$$\begin{pmatrix} \delta_{12} \\ \delta_{13} \\ \vdots \\ \delta_{1p} \end{pmatrix}_{(p-1) \times 1} = \begin{pmatrix} \delta_{22} & \delta_{23} & \dots & \delta_{2j} & \dots & \delta_{2p} \\ \delta_{23} & \delta_{33} & & \delta_{3j} & & \delta_{3p} \\ \vdots & \vdots & & \vdots & & \vdots \\ \delta_{2p} & \delta_{3p} & \dots & \delta_{jp} & \dots & \delta_{pp} \end{pmatrix}_{(p-1) \times (p-1)} \begin{pmatrix} b_2 \\ b_3 \\ \vdots \\ b_p \end{pmatrix}_{(p-1) \times 1}$$

By Cramer's rule

$$b_2 = \frac{
 \begin{vmatrix} \delta_{12} & \delta_{23} & \dots & \delta_{2p} \\ \delta_{13} & \delta_{33} & \dots & \delta_{3p} \\ \vdots & \vdots & & \vdots \\ \delta_{1p} & \delta_{3p} & \dots & \delta_{pp} \end{vmatrix}
 }{
 \begin{vmatrix} \delta_{22} & \delta_{23} & \dots & \delta_{2p} \\ \delta_{23} & \delta_{33} & \dots & \delta_{3p} \\ \vdots & \vdots & & \vdots \\ \delta_{2p} & \delta_{3p} & \dots & \delta_{pp} \end{vmatrix}
 }$$

In general,

$$b_j = \frac{
 \begin{vmatrix} \delta_{22} & \delta_{23} & \dots & \delta_{2,j-1} & \delta_{12} & \delta_{2,j+1} & \dots & \delta_{2p} \\ \delta_{23} & \delta_{33} & & \delta_{3,j-1} & \delta_{13} & \delta_{3,j+1} & \dots & \delta_{3p} \\ \vdots & \vdots & & \vdots & \vdots & \vdots & & \vdots \\ \delta_{2p} & \delta_{3p} & \dots & \delta_{p,j-1} & \delta_{1p} & \delta_{p,j+1} & \dots & \delta_{pp} \end{vmatrix}
 }{
 \begin{vmatrix} \delta_{22} & \delta_{23} & \dots & \delta_{2p} \\ \delta_{23} & \delta_{33} & \dots & \delta_{3p} \\ \vdots & \vdots & & \vdots \\ \delta_{2p} & \delta_{3p} & \dots & \delta_{pp} \end{vmatrix}
 }$$



$$= (-1)^{j-2} \begin{vmatrix} \delta_{12} & \delta_{22} & \delta_{23} & \dots & \delta_{2,j-1} & \delta_{2,j+1} & \dots & \delta_{2p} \\ \delta_{13} & \delta_{23} & \delta_{33} & \dots & \delta_{3,j-1} & \delta_{3,j+1} & \dots & \delta_{3p} \\ \vdots & \vdots & \vdots & \ddots & \vdots & \vdots & \ddots & \vdots \\ \delta_{1p} & \delta_{2p} & \delta_{3p} & \dots & \delta_{p,j-1} & \delta_{p,j+1} & \dots & \delta_{pp} \end{vmatrix}$$

$$\begin{vmatrix} \delta_{22} & \delta_{23} & \dots & \delta_{2p} \\ \delta_{23} & \delta_{33} & \dots & \delta_{3p} \\ \vdots & \vdots & \ddots & \vdots \\ \delta_{2p} & \delta_{3p} & \dots & \delta_{pp} \end{vmatrix}$$

$$= \left\langle \frac{(-1)^{j-2} \cdot (-1)^{-(1+j)} \text{cofactor of } \delta_{1j}}{(-1)^{-(1+1)} \text{cofactor of } \delta_{11}} \right\rangle$$

[cofactor of  $(i, j)$  element =  $(-1)^{i+j}$  minor of  $(i, j)$  element]

minors of  $(i, j)$  element =  $(-1)^{-i+j}$  cofactor of  $(i, j)$  element]

$$\Rightarrow b_j = - \frac{S_{i,j}}{S_{1,1 \dots i}}$$

where,  $S_{i,j}$  being the cofactor of  $(i, j)$  element in  $S$ .

Another representation of  $b_j$  comes from the correlation matrix  $R$

$$R = \begin{pmatrix} \text{corr}(x_1, x_1) & \text{corr}(x_1, x_2) & \dots & \dots & \text{corr}(x_1, x_p) \\ \text{corr}(x_2, x_1) & \text{corr}(x_2, x_2) & \dots & \dots & \text{corr}(x_2, x_p) \\ \vdots & \vdots & \ddots & \vdots & \vdots \\ \text{corr}(x_p, x_1) & \text{corr}(x_p, x_2) & \dots & \dots & \text{corr}(x_p, x_p) \end{pmatrix}$$

\$p \times p\$

$$= \begin{pmatrix} 1 & r_{12} & r_{13} & \dots & r_{1p} \\ r_{12} & 1 & r_{23} & \dots & r_{2p} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ r_{p1} & r_{p2} & \dots & \dots & 1 \end{pmatrix}$$

as \$r\_{11} = r\_{22} = \dots = r\_{pp} = 1\$

$$\therefore |R| = \begin{vmatrix} 1 & r_{12} & r_{13} & \dots & r_{1p} \\ r_{12} & 1 & r_{23} & \dots & r_{2p} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ r_{p1} & r_{p2} & \dots & \dots & 1 \end{vmatrix}$$

$$|R| = r_{11} R_{1,1} + r_{12} R_{1,2} + \dots + r_{1p} R_{1,p}$$

$$b_j = (-1)^{j-2} \begin{vmatrix} r_{12} \delta_1 \delta_2 & r_{22} \delta_2^2 & \dots & r_{1,j-1} \delta_2 \delta_{j-1} & \dots & r_{1p} \delta_2 \delta_p \\ \vdots & \vdots & \ddots & \vdots & \ddots & \vdots \\ r_{1p} \delta_1 \delta_p & r_{2p} \delta_2 \delta_p & \dots & r_{p,j-1} \delta_p \delta_{j-1} & \dots & r_{pp} \delta_p^2 \end{vmatrix}$$

$$\begin{vmatrix} r_{22} \delta_2^2 & r_{23} \delta_2 \delta_3 & \dots & r_{2p} \delta_2 \delta_p \\ \vdots & \vdots & \ddots & \vdots \\ r_{2p} \delta_2 \delta_p & r_{3p} \delta_3 \delta_p & \dots & r_{pp} \delta_p \delta_p \end{vmatrix}$$

$$= (-1)^{j-2} \delta_1^2 \delta_2^2 \delta_3^2 \dots \delta_{j-1}^2 \delta_{j+1} \delta_j \dots \delta_p^2 \begin{vmatrix} r_{12} r_{22} & r_{23} r_{2,j-1} & \dots & r_{2p} \\ \vdots & \vdots & \ddots & \vdots \\ r_{1p} r_{2p} & r_{p,j-1} & \dots & r_{pp} \end{vmatrix}$$

$$= \begin{vmatrix} \delta_2^2 \delta_3^2 \dots \delta_j^2 \dots \delta_p^2 & r_{22} & r_{23} & \dots & r_{2p} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ r_{2p} r_{3p} & r_{3p} & \dots & r_{pp} \end{vmatrix}$$

$$\Rightarrow b_j = \frac{(-1)^{j-2} \delta_1 \cdot (-1)^{-(1+j)} \text{cofactor of } (1, j) \text{ element in } R}{(-1)^{-(1+1)} \text{cofactor of } (1, 1) \text{ element in } R}$$

$$= \frac{-\delta_1}{\delta_j} \frac{\text{Cofactor of } (1, j) \text{ element in } R}{\text{Cofactor of } (1, 1) \text{ element in } R}$$

$$= -\frac{\delta_1}{\delta_j} \frac{R_{1,j}}{R_{1,1}} = -\frac{R_{1,j}}{R_{1,1}} \cdot \frac{\delta_1}{\delta_j}$$

< where  $R_{i,j}$  being the cofactor of  $(i, j)$  element in  $R$  >

Finally the multiple regression line

$$X_1 = a + b_2 X_2 + b_3 X_3 + \dots + b_p X_p$$

$$= \bar{X}_1 - b_2 \bar{X}_2 - b_3 \bar{X}_3 - \dots - b_p \bar{X}_p + b_2 X_2 + b_3 X_3 + \dots + b_p X_p$$

$$= \bar{X}_1 + b_2 (X_2 - \bar{X}_2) + b_3 (X_3 - \bar{X}_3) + \dots + b_p (X_p - \bar{X}_p)$$

$$= \bar{X}_1 - \frac{R_{1,2}}{R_{1,1}} \cdot \frac{\delta_1}{\delta_2} (X_2 - \bar{X}_2) - \frac{R_{1,3}}{R_{1,1}} \cdot \frac{\delta_1}{\delta_3} (X_3 - \bar{X}_3) - \dots - \frac{R_{1,p}}{R_{1,1}} \cdot \frac{\delta_1}{\delta_p} (X_p - \bar{X}_p)$$

Multiple regression line of  $x_1$  on  $x_2, x_3, \dots, x_p$

$$\therefore \langle x_1 = x_{1.23\dots p} + e_{1.23\dots p} \rangle$$

$$\left\{ \begin{aligned} x_{1.23\dots p} &= \bar{x}_1 - \frac{R_{1,2}}{R_{1,1}} \frac{s_1}{s_2} (x_2 - \bar{x}_2) - \frac{R_{1,3}}{R_{1,1}} \frac{s_1}{s_3} \\ &\quad (x_3 - \bar{x}_3) - \dots - \frac{R_{1,p}}{R_{1,1}} \frac{s_1}{s_p} (x_p - \bar{x}_p) \end{aligned} \right.$$

In multiple regression line  $b_j$ 's are called partial regression coefficient and mathematically denoted by

$$\boxed{b_{j.23\dots j-1; j+1\dots p}}$$

where

$$\left\langle b_{j.23\dots j-1; j+1\dots p} = - \frac{R_{1,j}}{R_{1,1}} \frac{s_1}{s_j} \right\rangle$$

► We write it in a short format as  $b_j$

Why Partial?

•  $b_j$  measures the amount of change in dependent variable  $x_1$  while unit amount of change occurred in independent variable  $x_j$  keeping the other ind. variables  $x_2, x_3, \dots, x_{j-1}, x_{j+1}, \dots, x_p$  fixed.

## Remark

$$(-\infty < b_j < \infty)$$

$$x_1 = 1 - 2.2x_2 + 3.9x_3 - 4x_4 \Rightarrow \text{Interpretation}$$

There has a inverse relation b/w

$$(x_1, x_2) \text{ and } (x_1, x_4)$$

< If  $x_4$  dec by 1 unit then  $x_1$  will inc. by 4 unit

If  $x_3$  inc. by 1 :  $x_1$  inc by 3.9 unit >

① Predicted value of  $x_1$  at  $\alpha$ th observation

$$x_{1\alpha} = \bar{x}_1 - \frac{R_{1,2}}{R_{1,1}} \frac{s_1}{s_2} (x_{2\alpha} - \bar{x}_2) - \frac{R_{1,3}}{R_{1,1}} \frac{s_1}{s_3} (x_{3\alpha} - \bar{x}_3)$$

$$- \frac{R_{1,p}}{R_{1,1}} \frac{s_1}{s_p} (x_{p\alpha} - \bar{x}_p)$$

② Multiple Correlation coefficient :

Def >> To measure the degree of association between the dependent variable  $x_1$  and the set of independent variables  $(x_2, x_3, \dots, x_p)$  and extension of Pearson correlation coefficient is introduced and that is called multiple correlation coefficient.

(How effective (meaningful) the linear regression in explaining  $x_1$  through  $x_2, x_3, \dots, x_p$ )

It is denoted by  $r_{1.23\dots p}$  and  $r$ .

$$r_{1.23\dots p} = \text{Corr}(x_1, x_{1.23\dots p})$$

$$= \frac{\text{Cov}(x_1, x_{1.23\dots p})}{\sqrt{V(x_1) V(x_{1.23\dots p})}}$$

Derivation

Least square multiple regression line is

$$x_1 = \bar{x}_1 - \frac{R_{1,2}}{R_{1,1}} \frac{\delta_1}{\delta_2} (x_2 - \bar{x}_2) - \dots - \frac{R_{1,p}}{R_{1,1}} \frac{\delta_1}{\delta_p} (x_p - \bar{x}_p)$$

$$V(x_1) = \delta_{11} = \delta_1^2$$

The model is

$$x_1 = x_{1.23\dots p} + e_{1.23\dots p}$$

$$x_{1.23\dots p} = \bar{x}_1 - \frac{R_{1,2}}{R_{1,1}} \frac{\delta_1}{\delta_2} (x_2 - \bar{x}_2) - \dots - \frac{R_{1,p}}{R_{1,1}} \frac{\delta_1}{\delta_p} (x_p - \bar{x}_p)$$

$$\bar{x}_{1.23\dots p} = \frac{1}{n} \sum_{\alpha=1}^n x_{1.23\dots p}$$

$$= \frac{1}{n} \sum_{\alpha=1}^n \left( \bar{x}_1 - \frac{R_{1,2}}{R_{1,1}} \frac{\delta_1}{\delta_2} (x_{2\alpha} - \bar{x}_2) - \dots - \frac{R_{1,p}}{R_{1,1}} \frac{\delta_1}{\delta_p} (x_{p\alpha} - \bar{x}_p) \right)$$

$$= \bar{x}_1 - 0 - \dots - 0 = \bar{x}_1$$

$$\text{Also, } \text{Cov}(x_1, x_{1.23\dots p})$$

$$= \text{Cov}(x_{1.23\dots p} + e_{1.23\dots p}, x_{1.23\dots p})$$

$$= V(x_{1.23\dots p}) + \text{Cov}(e_{1.23\dots p}, x_{1.23\dots p})$$

$$\text{Cov}(x+y, z)$$

$$= \text{Cov}(x, z)$$

$$+ \text{Cov}(y, z)$$

$$\text{Now, } \text{Cov}(e_{1.23\dots p}, x_{1.23\dots p})$$

$$= \frac{1}{n} \sum_{\alpha} (e_{1.23\dots p\alpha} - \bar{e}_{1.23\dots p}) (x_{1.23\dots p\alpha} - \bar{x}_{1.23\dots p})$$

$$= \frac{1}{n} \sum_{\alpha} e_{1.23\dots p\alpha} (a_1 + b_2 x_{2\alpha} + \dots + b_p x_{p\alpha} - \bar{x}_1) \quad \left[ \because \bar{e}_{1.23\dots p} = 0 \right]$$

(1st normal eqn)

$$= \frac{1}{n} \sum_{\alpha} e_{1.23\dots p\alpha} (a_1 \bar{x}_1 - b_2 \bar{x}_2 - \dots - b_p \bar{x}_p + b_2 x_{2\alpha} + b_3 x_{3\alpha} + \dots + b_p x_{p\alpha} - \bar{x}_1)$$

$$= \frac{1}{n} \left[ \sum_{\alpha} b_2 e_{1.23\dots p\alpha} (x_{2\alpha} - \bar{x}_2) + b_3 \sum_{\alpha} e_{1.23\dots p\alpha} (x_{3\alpha} - \bar{x}_3) + \dots + b_p \sum_{\alpha} e_{1.23\dots p\alpha} (x_{p\alpha} - \bar{x}_p) \right]$$

$$= 0 \quad (\text{from normal equations})$$

$$\therefore r_{1.23\dots p} = \frac{V(x_{1.23\dots p})}{\sqrt{V(x_1) V(x_{1.23\dots p})}}$$

$$= \sqrt{\frac{V(x_{1.23\dots p})}{V(x_1)}}$$



$\therefore r_{1.23\dots p}$  is the ratio of two standard deviations, which leads to the range of

$$r_{1.23\dots p} \Rightarrow 0 < r_{1.23\dots p} < 1$$

$$x_1 = x_{1.23\dots p} + e_{1.23\dots p}$$

$$V(x_1) = V(x_{1.23\dots p}) + V(e_{1.23\dots p}) + 0$$

$$\therefore V(x_1) > V(x_{1.23\dots p})$$

In other words

$$r_{1.23\dots p} = \frac{\text{Cov}(x_1, x_{1.23\dots p})}{\sqrt{V(x_1) \text{Cov}(x_1, x_{1.23\dots p})}}$$

$$= \sqrt{\frac{\text{Cov}(x_1, x_{1.23\dots p})}{V(x_1)}}$$

$$\text{Cov}(x_1, x_{1.23\dots p})$$

$$= \frac{1}{n} \sum_{\alpha} (x_{1\alpha} - \bar{x}_1)(x_{1.23\dots p\alpha} - \bar{x}_{1.23\dots p})$$

$$= \frac{1}{n} \sum_{\alpha} (x_{1\alpha} - \bar{x}_1) \left( \bar{x}_1 - \frac{R_{1,2}}{R_{1,1}} \frac{\delta_1}{\delta_2} (x_{2\alpha} - \bar{x}_2) - \frac{R_{1,3}}{R_{1,1}} \frac{\delta_1}{\delta_3} (x_{3\alpha} - \bar{x}_3) \right)$$

$$- \dots - \frac{R_{1,p}}{R_{1,1}} \frac{\delta_1}{\delta_p} (x_{p\alpha} - \bar{x}_p) - \bar{x}_1$$

$$= -\frac{1}{n} \sum_{\alpha} (x_{1\alpha} - \bar{x}_1) (x_{2\alpha} - \bar{x}_2) \cdot \frac{R_{1,2}}{R_{1,1}} \frac{\delta_1}{\delta_2} - \frac{1}{n} \sum_{\alpha} (x_{1\alpha} - \bar{x}_1) (x_{3\alpha} - \bar{x}_3)$$

$$\frac{R_{1,3}}{R_{1,1}} \frac{\delta_1}{\delta_3} \dots - \frac{1}{n} \sum_{\alpha} (x_{1\alpha} - \bar{x}_1) (x_{p\alpha} - \bar{x}_p)$$

$$\frac{R_{1,p}}{R_{1,1}} \frac{\delta_1}{\delta_p}$$

$$= - \left( r_{12} \delta_1 \delta_2 \frac{R_{1,2}}{R_{1,1}} \frac{\delta_1}{\delta_2} + r_{13} \delta_1 \delta_3 \frac{R_{1,3}}{R_{1,1}} \frac{\delta_1}{\delta_3} + \dots + r_{1p} \delta_1 \delta_p \frac{R_{1,p}}{R_{1,1}} \frac{\delta_1}{\delta_p} \right) \left[ \frac{1}{n} \sum_k (x_{1k} - \bar{x}_1) \right]$$

$$= - \frac{\delta_1^2}{R_{1,1}} (r_{12} R_{1,2} + r_{13} R_{1,3} + \dots + r_{1p} R_{1,p}) \left[ \frac{1}{n} \sum_k (x_{2k} - \bar{x}_2) \right] = r_{12} \delta_1$$

$$= - \frac{\delta_1^2}{R_{1,1}} (r_{11} R_{1,1} + r_{12} R_{1,2} + r_{13} R_{1,3} + \dots + r_{1p} R_{1,p} - r_{11} R_{1,1})$$

$$= - \frac{\delta_1^2}{R_{1,1}} (|R| - R_{1,1}) \quad \left[ \begin{array}{l} \because r_{11} = 1 \\ = \text{Cov}(x_1, x_1) \end{array} \right]$$

$$= \delta_1^2 \left( 1 - \frac{|R|}{R_{1,1}} \right)$$

$$\therefore r_{1,23\dots p} = \frac{\text{Cov}(x_1, x_{1,23\dots p})}{\sqrt{V(x_1)}}$$

$$= \sqrt{\frac{\delta_1^2 \left( 1 - \frac{|R|}{R_{1,1}} \right)}{\delta_1^2}} = \left( 1 - \frac{|R|}{R_{1,1}} \right)^{1/2}$$

$$\therefore r_{2,13\dots p} = \left( 1 - \frac{|R|}{R_{2,2}} \right)^{1/2} \quad \left[ \begin{array}{l} |R| < 1 \\ \text{positive} \\ \text{definite} \\ \text{matrix} \end{array} \right]$$

Prop 1 :  $r_{1.23\dots p} = 1$

$\Rightarrow V(x_1) = V(x_{1.23\dots p})$

$\Rightarrow V(e_{1.23\dots p}) = 0$

$\Rightarrow$  There is no variability of  $e_{1.23\dots p}$

$\Rightarrow$  Error is constant

$\Rightarrow$  The prediction formula is perfect.  
(linear)

▣  $r_{1.23\dots p}^2 = 0$

$\Rightarrow V(x_{1.23\dots p}) = 0$

$\Rightarrow x_{1.23\dots p} = a + b_2 x_2 + \dots + b_p x_p$  is independent of  $x_2, x_3, \dots, x_p$

$\Rightarrow$  The prediction formula fails to predict the dependent variable  $x_1$ .

Remark

▣  $r_{1.23\dots p}^2 = \frac{V(x_{1.23\dots p})}{V(x_1)}$

$r_{1.23\dots p}^2$  is called coefficient of determination and generally denoted by  $R^2$ .

Coefficient of determination measures the variance of dependent variable explained by the prediction or regression equation,  $x_1 = a + b_2 x_2 + \dots + b_p x_p$

$R^2$  is generally denoted ~~by~~ in percentage format.

For example,  $r^2_{1.23-p} = 0.35$

$$\Rightarrow R^2 = 35\%$$

35% of the total variability of dependent variable ( $X_1$ ) is explained (determined) by the multiple regression line

$$X_1 = a + b_2 X_2 + \dots + b_p X_p$$

65% of the variability of  $X_1$  remains unexplained

Remark

With the increased of independent variables in the model the value of  $R^2$  also increases.

$$0 < R^2 < 1$$

$$X_1 = a + b_2 X_2$$

$$R_1^2$$

$$X_2 = a + b_2 X_2 + b_3 X_3$$

$$R_2^2 > R_1^2$$

$$X_1 = a + b_2 X_2 + \dots + b_p X_p$$

$$R_p^2 > R_2^2 > R_1^2$$

$$\boxed{\therefore R_p^2 > R_{p-1}^2 > R_{p-2}^2 > \dots > R_2^2 > R_1^2}$$

Now to check on the value of  $R^2$ , generally we consider upto 95% of explainability.

- 60% of variable might have check whether they are present on your model or not.

Later on there discovered a modification of  $R^2$  which is called adjusted  $R^2$ .

Adjusted  $R^2$  will not change its value significantly even if inclusion of new ~~variable~~ independent variable in the model.

Remark: ~~Standard error of estimate~~  
(residual variance)

$$V(e_{1,23 \dots P})$$

$$= V(x_1) - V(x_{1,23 \dots P})$$

$$= \delta_1^2 - \text{Cov}(x_{1,23 \dots P}, x_1)$$

$$= \delta_1^2 - \delta_1^2 \left(1 - \frac{|R|}{R_{1,1}}\right)$$

$$= \delta_1^2 \frac{|R|}{R_{1,1}}$$

$$\text{Standard error} = \sqrt{\delta_1^2 \frac{|R|}{R_{1,1}}}$$

Exercise 1: Multiple Regression,  
Multiple and Partial Correlation

Pmb ① The following data relates to the ~~person~~ performance of the students in the final examination ( $x_1$ ), 1st term examination ( $x_2$ ), 2nd term examination ( $x_3$ ), 3rd term examination ( $x_4$ )

(a) Fit a <sup>least square</sup> regression equation for  $x_1$  on  $x_2, x_3, x_4$  as ①  $x_1 = a + b_2 x_2 + b_3 x_3 + b_4 x_4$

② when  $b_2 = b_3 = b_4$  and ③  $b_2 + b_3 + b_4 = 1$   
(restricted domain regression)

**Hint**

①  $x_1 = a + b_2(x_2 + x_3 + x_4)$   
 $\Rightarrow a + b_2 z$

③  $x_1 = a + (1 - b_3 - b_4)x_2 + b_3 x_3 + b_4 x_4$

② Obtain the multiple corr. coeff. <sup>of</sup>  $x_1$  on  $x_2, x_3, x_4$

Student Roll No	$X_2$	$X_3$	$X_4$	$X_1$
1	67.8	71.2	60.0	58.6
2	68.7	67.0	73.0	55.5
3	71.2	73.0	69.0	67.1
4	65.0	77.5	60.5	54.7
5	53.4	75.0	57.0	43.2
6	70.8	78.5	43.0	60.1
7	63.4	71.0	55.5	52.7
8	67.4	69.0	61.5	51.7
9	79.9	87.0	46	66.5
10	66.5	72.5	52.5	50.4
11	82.2	85.0	57.5	61.1
12	69.6	70.5	47.0	68.0
13	72.2	81	47.5	50.7
14	69.1	82.5	47.5	65.6
15	60.5	71.5	51.0	46.2

$$R = \begin{pmatrix} 1 & r_{12} & r_{13} & r_{14} \\ r_{12} & 1 & r_{23} & r_{24} \\ r_{31} & r_{32} & 1 & r_{34} \\ r_{41} & r_{42} & r_{43} & 1 \end{pmatrix}$$

$$r_{12} = \frac{\text{Cov}(x_1, x_2)}{\sqrt{v(x_1) \cdot v(x_2)}}$$

① Mean

$$\bar{x}_1 = 56.806$$

$$\bar{x}_2 = 68.513$$

$$\bar{x}_3 = 75.48$$

$$\bar{x}_4 = 55.23$$

$$\text{Cov}(x_1, x_2)$$

$$= 3927.696 - (56.806 \times 68.513)$$

$$= 35.746$$

$$r_{12} = \frac{35.746}{\dots}$$

$$r_{12} = 0.7$$

$$r_{13} = 0.339$$

$$r_{14} = -0.114$$

$$r_{23} = 0.579$$

$$r_{24} = 0.342$$

$$r_{34} = -0.52$$

② Var

$$\text{var}(x_1) = \dots$$

$$= 57.335$$

$$s_2 = 6.74$$

$$\text{var}(x_2) = 45.433$$

$$s_3 = 5.913$$

$$\text{var}(x_3) = 34.965$$

$$s_4 = 8.372$$

$$\text{var}(x_4) = 70.099$$

$$\text{Cov}(x_1, x_3)$$

$$= 15.1808$$

$$\text{Cov}(x_1, x_4)$$

$$= -7.0644$$

$$\text{Cov}(x_2, x_3) = 23.1125$$

$$\text{Cov}(x_2, x_4)$$

$$= 19.296$$

$$\text{Cov}(x_3, x_4)$$

$$= -25.744$$



$$R = \begin{pmatrix} 1 & 0.7 & 0.339 & -0.114 \\ 0.7 & 1 & 0.579 & 0.342 \\ 0.339 & 0.579 & 1 & -0.52 \\ -0.114 & 0.342 & -0.52 & 1 \end{pmatrix}$$

$R_{1,3} = \frac{R_{1,2} R_{1,3}}{R_{1,1} R_{1,1} R_{1,3}}$

$$R_{1,3} = (-1)^{1+3} \begin{pmatrix} 0.7 & 1 & 0.342 \\ 0.339 & 0.579 & -0.52 \\ -0.114 & 0.342 & 1 \end{pmatrix}$$

$$= -(-0.529 - 0.279 - 0.062)$$

$$= 0.188 \quad R_{1,2} = 0.1823$$

$$R_{1,1} = 0.211$$

$$R_{1,4} = (-1)^{1+4} \begin{pmatrix} 0.7 & 1 & 0.579 \\ 0.339 & 0.579 & 1 \\ -0.114 & 0.342 & -0.52 \end{pmatrix}$$

$$= -[-0.45 + 0.062 + 0.105] = 0.283$$

$$x_1 = \bar{x}_1 - \frac{R_{1,2}}{R_{1,1}} \frac{s_1}{s_2} (x_2 - \bar{x}_2) - \frac{R_{1,3}}{R_{1,1}} \frac{s_1}{s_3} (x_3 - \bar{x}_3) - \frac{R_{1,4}}{R_{1,1}} \frac{s_1}{s_4} (x_4 - \bar{x}_4)$$

$$= 56.806 - \frac{0.182}{0.211} \times \frac{7.572}{6.74} (x_2 - 68.513)$$

$$- \frac{0.188}{0.211} \times \frac{7.572}{5.913} (x_3 - 75.48) - \frac{0.283}{0.211} \times \frac{7.572}{8.372} (x_4 - 55)$$

$$= 56.806 - 0.97x_2 + 66.392 - 1.41x_3 + 86.121 - 1.213x_4 + 66.9$$

$$\boxed{x_1 = 276.316 - 0.97x_2 - 1.41x_3 - 1.213x_4}$$

$$\boxed{x_1 = 17.724 + 0.88x_2 - 0.213x_3 - 0.001x_4}$$

multiple corr. coeff.

$$r_{1,2,3,4} = \left(1 - \frac{|R|}{R_{1,1}}\right)^{1/2}$$

$$|R| = 1 \times 0.211 + (0.4 \times 0.1823) + (0.339 \times 0.188) + (0.114 \times \dots)$$

$$= 0.37098$$

$$r_{1,2,3,4} = \left(1 - \frac{0.37098}{0.211}\right)^{1/2}$$

$$|R| = 0.2280337, \quad R_{1,1} = 0.4495$$

$$r_{1,2,3,4} = 0.709$$

$$R = \begin{pmatrix} 1 & 0.699 & 0.339 & -0.117 \\ 0.699 & 1 & 0.586 & -0.143 \\ 0.339 & 0.586 & 1 & -0.525 \\ -0.117 & -0.143 & -0.525 & 1 \end{pmatrix}$$

$$R_{1,1} = 0.4485, \quad R_{1,2} = -0.3524$$

$$R_{1,3} = 0.0767$$

$$R_{1,4} = 0.0423$$

$$b_3 = - \frac{0.0767}{0.4485} \times \frac{7.572}{5.913} = -0.2189$$

$$b_4 = - \frac{0.0423}{0.4485} \times \frac{7.572}{8.372} = -0.0853$$

.23)

97

8674

-0.283)

Q.1. For three variables  $x_1, x_2, x_3$  prove that  

$$r_{12}^2 + r_{13}^2 + r_{23}^2 - 2r_{12}r_{13}r_{23} \leq 1$$

Ans:- Model  $x_1 = a + b_2x_2 + b_3x_3$   

$$= \bar{x}_1 - \frac{R_{1,2}}{R_{1,1}} \frac{b_1}{b_2} (x_2 - \bar{x}_2) - \frac{R_{1,3}}{R_{1,1}} \frac{b_1}{b_3} (x_3 - \bar{x}_3)$$

where,  $R = \begin{pmatrix} 1 & r_{12} & r_{13} \\ r_{12} & 1 & r_{23} \\ r_{13} & r_{23} & 1 \end{pmatrix}$

~~where~~  $r_{1.23} = \left( 1 - \frac{|R|}{R_{1,1}} \right)^{1/2}$

$$\begin{aligned} |R| &= 1 \cdot (1 - r_{23}^2) - r_{12}(r_{12} - r_{13}r_{23}) + r_{13}(r_{12}r_{23} - r_{13}) \\ &= 1 - r_{23}^2 - r_{12}^2 + r_{12}r_{13}r_{23} + r_{12}r_{13}r_{23} - r_{13}^2 \\ &= 1 - r_{23}^2 - r_{12}^2 + 2r_{12}r_{13}r_{23} - r_{13}^2 \end{aligned}$$

$$R_{1,1} = 1 - r_{23}^2$$

$$r_{1.23} = \left( 1 - \frac{1 - r_{23}^2 - r_{12}^2 + 2r_{12}r_{13}r_{23} - r_{13}^2}{1 - r_{23}^2} \right)^{1/2}$$

$$r_{1.23}^2 \leq 1$$

$$\frac{1 - r_{23}^2 - 1 + r_{23}^2 + r_{12}^2 - 2r_{12}r_{13}r_{23} + r_{13}^2}{1 - r_{23}^2} \leq 1$$

$$r_{12}^2 + r_{13}^2 + r_{23}^2 - 2r_{12}r_{13}r_{23} \leq 1$$

Prob 2

Suppose for a regression model,  $x_1, x_2, \dots, x_p$

Correlation of  $x_i, x_j$ ,  $\text{Corr}(x_i, x_j) = r_{ij}$

$i=1(1)P \quad i \neq j$   
 $j=1(1)P$

Find  $r_{1.23 \dots p}$

Ans:

$$R = \begin{pmatrix} 1 & r_{12} & r_{13} & \dots & r_{1p} \\ r_{12} & 1 & r_{23} & \dots & r_{2p} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ r_{1p} & r_{2p} & r_{3p} & \dots & 1 \end{pmatrix}_{p \times p}$$

Intraclass pattern matrix.

$$|R| = (1-r)^{p-1} (1 + (p-1)r)$$

since  $|R| > 0$

$$R = (1-r)^{p-1} (1 + (p-1)r)$$

$$M = \begin{pmatrix} a & b & b & \dots & b \\ b & a & b & \dots & b \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ b & b & \dots & \dots & a \end{pmatrix}_{n \times n}$$

$$|M| = (b-a)^{n-1} (a + (n-1)b)$$

$$= (a-b)^{n-1} (a + (n-1)b)$$

$$R_{1,1} = (1-r)^{P-2} (1 + \overline{P-2} \cdot r)$$

$$r_{1.23} \dots P = \left( \frac{(1-r)^{P-1} (1 + \overline{P-1} \cdot r)}{(1-r)^{P-2} (1 + \overline{P-2} \cdot r)} \right)^{1/2}$$

$$= \left( \frac{(1-r)^{P-2} (1 + \overline{P-2} \cdot r) - (1-r)^{P-1} (1 + \overline{P-1} \cdot r)}{(1-r)^{P-2} (1 + \overline{P-2} \cdot r)} \right)^{1/2}$$

$$= \left( \frac{(1-r)^{P-2} \cdot [1 + \overline{P-2} \cdot r - (1-r)(1 + \overline{P-1} \cdot r)]}{(1-r)^{P-2} (1 + \overline{P-2} \cdot r)} \right)^{1/2}$$

$$= \left( \frac{\cancel{(P-2)r} :}{(1-r)^{P-2} (1 + \overline{P-2} \cdot r)} \right)^{1/2}$$

$$= \left( \frac{1 + \cancel{r(P-2)r} - 1 + \cancel{P}r + \cancel{r} + \cancel{r^2 P} + \cancel{r^2}}{1 + (P-2)r} \right)^{1/2}$$

$$= \left( \frac{r^2 \cdot (P-1)}{1 + (P-2)r} \right)^{1/2}$$

$$= r \left( \frac{P-1}{1 + (P-2)r} \right)^{1/2}$$

Prob 3

Suppose for a  $p$ -dimensional model

$$r_{ij} = r, \quad i, j = 1, \dots, p$$

$$r_{ij} = r', \quad i = 1, \dots, p, \quad j = 1, \dots, p$$

Find  $r_{1,2,3, \dots, p}$

$$R_{p \times p} = \left( \begin{array}{c|cccc} \text{(A)} & & & & \\ \hline & r & r & \dots & r \\ \hline & r & 1 & r' & \dots & r' \\ & r & r' & 1 & \dots & r' \\ \vdots & \vdots & \vdots & \vdots & \ddots & \vdots \\ & r & r' & r' & \dots & 1 \text{ (D)} \end{array} \right)_{p \times p}$$

$$R_{1,1} = (1-r')^{p-2} (1 + (p-2) \cdot r')$$

~~$|R| = |I|$~~

partition determinant

$$M = \begin{pmatrix} A & B \\ C & D \end{pmatrix}$$

$$|M| = |A| |D - CA^{-1}B|$$

$$= |D| |A - BD^{-1}C|$$

$$|R| = |I| \left| \begin{array}{c} 1 \quad r' \dots r' \\ r' \quad 1 \dots r' \\ \vdots \\ r' \quad r' \dots 1 \end{array} \right| - r \cdot 1 \cdot r'$$

$$\underline{r}' = (r \ r \dots r)$$

$$\underline{r} = \begin{pmatrix} r \\ r \\ \vdots \\ r \end{pmatrix}$$

$$= \left| \begin{array}{c} 1 \quad r' \dots r' \\ r' \quad 1 \dots r' \\ \vdots \\ r' \quad r' \dots 1 \end{array} \right| - \left| \begin{array}{c} r^2 \quad r^2 \dots r^2 \\ r^2 \quad r^2 \dots r^2 \\ \vdots \\ r^2 \quad r^2 \dots r^2 \end{array} \right|$$

$$= \begin{pmatrix} 1-r^2 & r'-r^2 & r'-r^2 & \dots & r'-r^2 \\ r'-r^2 & 1-r^2 & r'-r^2 & \dots & r'-r^2 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ r'-r^2 & r'-r^2 & r'-r^2 & \dots & 1 \end{pmatrix}$$

~~P~~ (P-1) x P

$$a = 1-r^2$$

$$b = r'-r^2$$

$$= (1-r^2 - r' + r^2)^{P-2} \left( \frac{1}{1-r^2} (P-2) (r'-r^2) \right)$$

$$= (1-r')^{P-2} \left[ 1-r^2 + Pr' - Pr^2 - 2r' + 2r^2 \right]$$

$$\rightarrow (1-r')^{P-2} \left[ 1 + (P-2)r' - (P-1)r^2 \right]$$

$$r_{1.23 \dots P} = \left( 1 - \frac{1 + (P-2)r' - (P-1)r^2}{1 + (P-2)r'} \right)^{1/2}$$

$$\rightarrow r \left( \frac{(P-1)}{1 + (P-2)r'} \right)^{1/2}$$



## Partial Correlation coefficient :

$$(x_1, x_2, \dots, x_p)$$

$$x_1 = a + b_2 x_2 + \dots + b_p x_p$$

$$x_1 = \bar{x}_1 - \frac{R_{1,2}}{R_{1,1}} \frac{s_1}{s_2} (x_2 - \bar{x}_2) - \dots - \frac{R_{1,p}}{R_{1,1}} \frac{s_1}{s_p} (x_p - \bar{x}_p)$$

$x_1$  vs  $x_2, \dots, x_p$

$$x_1 \rightarrow \begin{cases} x_2 \\ x_3 \\ \vdots \\ x_p \end{cases}$$

Def  $\Rightarrow$  Partial correlation is the measure of association between two variables in a multiple regression set up when influence of the other variables is eliminated.  
For a multiple regression setup  $(x_1, x_2, \dots, x_p)$  partial correlation between  $x_1$  and  $x_2$  is denoted by  $r_{12.34\dots p}$

Mathematically :

$$r_{12.34\dots p} = \frac{\text{Corr}(e_{1.34\dots p}, e_{2.34\dots p})}{\sqrt{V(e_{1.34\dots p}) V(e_{2.34\dots p})}}$$

$$= \frac{\text{Cov}(e_{1.34\dots p}, e_{2.34\dots p})}{\sqrt{V(e_{1.34\dots p}) V(e_{2.34\dots p})}}$$

$$\sqrt{V(e_{1.34\dots p}) V(e_{2.34\dots p})}$$

Why  $r_{12,34\dots p}$  is considered as a parameter

→ Think about the model

$$x_1 = a + b_2 x_2 + \dots + b_p x_p \quad (x_1, x_2, \dots, x_p)$$

(1) and  $x_2 = a^* + b_3^* x_3 + \dots + b_p^* x_p \quad (x_2, x_3, \dots, x_p)$

If we apply least square regression line,

$$R^{(2)} = \begin{pmatrix} r_{11} & r_{13} & \dots & r_{1p} \\ r_{21} & r_{23} & \dots & r_{2p} \\ \vdots & \vdots & \ddots & \vdots \\ r_{p1} & r_{p3} & \dots & r_{pp} \end{pmatrix}$$

← unexplained part of  $x_1$  by  $x_3, x_4, \dots, x_p$

$$x_1 = x_{1,34\dots p} + e_{1,34\dots p} \quad (\text{Model for the 1st eqn.})$$

↓ explained part of  $x_1$  by  $x_3, x_4, \dots, x_p$

$$x_2 = x_{2,34\dots p} + e_{2,34\dots p} \quad (\text{Model for 2nd eqn.})$$

↓ explained part of  $x_2$  by  $x_3, x_4, \dots, x_p$

← unexplained part of  $x_2$  by  $x_3, x_4, \dots, x_p$

Derivation of  $r_{12,34\dots p}$

Let the data structure

$$x_{11} \quad x_{12} \quad x_{13} \quad \dots \quad x_{1p}$$

$$x_{21} \quad x_{22} \quad x_{23} \quad \dots \quad x_{2p}$$

$$\vdots$$

$$x_{n1} \quad x_{n2} \quad x_{n3} \quad \dots \quad x_{np}$$

Multiple regression line of  $x_1, x_2, \dots, x_p$

$$x_1 = \bar{x}_1 - \frac{R_{1,2}}{R_{1,1}} \frac{\delta_1}{\delta_2} (x_2 - \bar{x}_2) - \frac{R_{1,3}}{R_{1,1}} \frac{\delta_1}{\delta_3} (x_3 - \bar{x}_3) - \dots - \frac{R_{1,p}}{R_{1,1}} \frac{\delta_1}{\delta_p} (x_p - \bar{x}_p)$$

Regression model is

$$x_1 = x_{1.23\dots p} + e_{1.23\dots p}$$

Let us think on two submodels

(i)  $x_1$  on  $(x_3, x_4, \dots, x_p)$

(ii)  $x_2$  on  $(x_3, x_4, \dots, x_p)$

Borrowing the idea of multiple regression model

(i)  $\Rightarrow x_1 = a^* + b_3^* x_3 + b_4^* x_4 + \dots + b_p^* x_p$   
 $x_1 = x_{1.34\dots p} + e_{1.34\dots p}$

(ii)  $\Rightarrow x_2 = a^0 + b_3^0 x_3 + b_4^0 x_4 + \dots + b_p^0 x_p$   
 $x_2 = x_{2.34\dots p} + e_{2.34\dots p}$

Proceeding the same way as we did in

multiple regression setup

$$x_{1.34\dots p} = \bar{x}_1 - \frac{R_{1,3}^{(1)}}{R_{1,1}^{(1)}} \frac{\delta_1}{\delta_3} (x_3 - \bar{x}_3)$$

$$- \frac{R_{1,p}^{(1)}}{R_{1,1}^{(1)}} \frac{\delta_1}{\delta_p} (x_p - \bar{x}_p)$$

$$x_{2.34\dots p} = \bar{x}_2 - \frac{R_{2,3}^{(2)}}{R_{2,2}^{(2)}} \frac{\delta_2}{\delta_3} (x_3 - \bar{x}_3)$$

$$- \frac{R_{2,p}^{(2)}}{R_{2,2}^{(2)}} \frac{\delta_2}{\delta_p} (x_p - \bar{x}_p)$$

where,

$$R = \begin{pmatrix} r_{11} & r_{12} & r_{13} & \dots & r_{1p} \\ r_{21} & r_{22} & r_{23} & \dots & r_{2p} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ r_{p1} & r_{p2} & r_{p3} & \dots & r_{pp} \end{pmatrix}$$

$R^{(1)}$  = Discarding 1st row and 1st column from

$$R^{(1)} = \begin{pmatrix} r_{22} & r_{23} & \dots & r_{2p} \\ r_{32} & r_{33} & \dots & r_{3p} \\ \vdots & \vdots & \ddots & \vdots \\ r_{p2} & r_{p3} & \dots & r_{pp} \end{pmatrix}$$

$R^{(2)}$  = Discarding 2nd row and 2nd column from  $R$

$$R^{(2)} = \begin{pmatrix} r_{11} & r_{13} & \dots & r_{1p} \\ r_{31} & r_{33} & \dots & r_{3p} \\ \vdots & \vdots & \ddots & \vdots \\ r_{p1} & r_{p3} & \dots & r_{pp} \end{pmatrix}$$

$R_{1,3}^{(2)}$  = Cofactor of  $r_{13}$  in  $R^{(2)}$

$R_{ij}^{(k)}$  = Cofactor of  $r_{ij}$  in  $R^{(k)}$ ,  $k=1, 2, \dots$

Also we know,

$$\text{Cov}(x_1, x_{1.34 \dots p}) = V(x_{1.34 \dots p})$$

$$\text{and } \text{Cov}(x_2, x_{2.34 \dots p}) = V(x_{2.34 \dots p})$$

Again,

$$\begin{aligned} & \text{Cov}(x_1, x_{1.34 \dots p}) \\ &= \frac{1}{n} \sum_{\alpha=1}^n (x_{1\alpha} - \bar{x}_1) (x_{1.34 \dots p\alpha} - \bar{x}_{1.34 \dots p}) \end{aligned}$$

$$\begin{aligned} &= \frac{1}{n} \sum_{\alpha=1}^n (x_{1\alpha} - \bar{x}_1) \left( x_{1\alpha} - \frac{R_{1,3}^{(2)}}{R_{1,1}^{(2)}} \frac{\delta_1}{\delta_3} (x_{3\alpha} - \bar{x}_3) \right. \\ & \quad \left. - \dots - \frac{R_{1,p}^{(2)}}{R_{1,1}^{(2)}} \frac{\delta_1}{\delta_p} (x_{p\alpha} - \bar{x}_p) - \bar{x}_1 \right) \end{aligned}$$

$$\begin{aligned} & \bar{x}_{1.23 \dots p} \\ &= \bar{x}_1 \\ & \frac{1}{n} \sum_{\alpha} (x_{1\alpha} - \bar{x}_1) \\ & \quad (x_{3\alpha} - \bar{x}_3) \\ &= r_{13} \delta_1 \delta_3 \end{aligned}$$

$$= - \frac{R_{1,3}^{(2)}}{R_{1,1}^{(2)}} r_{13} \delta_1 \delta_3 \cdot \frac{\delta_1}{\delta_3} - \frac{R_{1,4}^{(2)}}{R_{1,1}^{(2)}} r_{14} \delta_1 \delta_4 \cdot \frac{\delta_1}{\delta_4}$$

$$- \dots - \frac{R_{1,p}^{(2)}}{R_{1,1}^{(2)}} \frac{\delta_1}{\delta_p} r_{1p} \delta_1 \delta_p$$

$$= - \frac{\delta_1^2}{R_{1,1}^{(2)}} \left[ R_{1,3}^{(2)} r_{13} + \frac{R_{1,4}^{(2)} r_{14}}{\cancel{R_{1,1}^{(2)}}} + \dots + \frac{R_{1,p}^{(2)} r_{1p}}{\cancel{R_{1,1}^{(2)}}} \right]$$

$$\begin{aligned} R^{(2)} &= \begin{pmatrix} r_{11} & r_{13} & r_{14} & \dots & r_{1p} \\ r_{31} & r_{32} & r_{34} & \dots & r_{3p} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ r_{p1} & r_{p3} & \dots & \dots & r_{pp} \end{pmatrix} \\ &= - \frac{\delta_1^2}{R_{1,1}^{(2)}} [R - R_{1,1}^{(2)}] \\ &= \delta_1^2 \left( 1 - \frac{|R^{(2)}|}{R_{1,1}^{(2)}} \right) \\ &= V(x_{1.34 \dots p}) \end{aligned}$$

Similarly

$$V(\chi_{2.34 \dots p}) = \delta_2^2 \left( 1 - \frac{|R^{(1)}|}{R_{2,2}^{(1)}} \right)$$

$$\therefore V(\epsilon_{1.34 \dots p})$$

$$= V(\chi_1) - V(\chi_{1.34 \dots p})$$

$$= \delta_1^2 - \delta_1^2 \left( 1 - \frac{|R^{(2)}|}{R_{1,1}^{(2)}} \right)$$

$$= \delta_1^2 \frac{|R^{(2)}|}{R_{1,1}^{(2)}}$$

$$\left[ \begin{array}{l} \chi_1 = \chi_{1.34 \dots p} \\ + \epsilon_{1.34 \dots p} \\ V(\chi_1) = V(\chi_{1.34 \dots p}) \\ + V(\epsilon_{1.34 \dots p}) \end{array} \right]$$

$$\therefore V(\epsilon_{2.34 \dots p}) = V(\chi_2) - V(\chi_{2.34 \dots p})$$

$$= \delta_2^2 \frac{|R^{(1)}|}{R_{2,2}^{(1)}}$$

Prob 1

$$x_1 = a + b_2 x_2 + b_3 x_3 + b_4 x_4$$

✓ unrestricted multiple regression model

✓ multiple corr. coeff

○ Restricted regression model

(i)  $b_2 = b_3 = b_4$

$$x_1 = a + b_2 x_2 + b_2 x_3 + b_2 x_4$$

$$x_1 = a + b_2 (x_2 + x_3 + x_4)$$

$$x_1 = a + b_2 z$$

$$a = \bar{x}_1 - b_2 \bar{z}$$

$$b_2 = \frac{\text{Cov}(x_1, z)}{\text{Var}(x_1)}$$

	$x_1$	$z = x_2 + x_3 + x_4$
1	58.6	199
2	55	208.7
3		213.2
4		203
5		185.4
6		192.3
7		189.9
8		197.9
9		212.9
10		191.5
11		224.7
12		187.1
13		200.7
14		199.1
15		183

$$\bar{x}_1 = 56.806, \text{Var}(x_1) = 57.335$$

$$\bar{z} = 199.223$$

$$\begin{aligned} & \text{Cov}(x_1, x_2 + x_3 + x_4) \\ &= \text{Cov}(x_1, x_2) + \text{Cov}(x_1, x_3) \\ & \quad + \text{Cov}(x_1, x_4) \end{aligned}$$

$$= 43.8626$$

$$b_2 = 0.765$$

$$a = -95.599$$

$$\hat{x}_1 = -95.599 + 0.765z$$

$$(iii) b_2 + b_3 + b_4 = 1$$

$$\Rightarrow b_2 = 1 - b_3 - b_4$$

$$\hat{x}_1 = a + (1 - b_3 - b_4)\hat{x}_2 + b_3\hat{x}_3 + b_4\hat{x}_4$$

$$= a + \hat{x}_2 + b_3(\hat{x}_3 - \hat{x}_2) + b_4(\hat{x}_4 - \hat{x}_2)$$

~~$$\hat{x}_1 = a + \hat{x}_2 + b_3\hat{z}_1 + b_4\hat{z}_2$$~~

$$\hat{x}_1 = a + \hat{x}_2(1 - b_3 - b_4) + b_3\hat{x}_3$$

~~$$\hat{x}_1 - \hat{x}_2 = a - b_3\hat{z}_1 - b_4\hat{z}_2$$~~

$$a = \bar{x}_1 - \bar{x}_2 - b_3\bar{z}_1 - b_4\bar{z}_2$$

$$= \bar{x}_1 - \bar{x}_2 - b_3(\bar{x}_3 - \bar{x}_2) - b_4(\bar{x}_4 - \bar{x}_2)$$

$$a = -11.707 - b_3(6.967) - b_4(-13.283)$$

$$a = -11.707 - 6.967b_3 + 13.283b_4$$

~~$$b_3 = \frac{R_{1,3}}{R_{1,1}} \cdot \frac{\sqrt{V(x_1)}}{\sqrt{V(z_1)}}$$~~

~~$$b_4 = \frac{R_{1,4}}{R_{1,1}} \cdot \frac{\sqrt{V(x_1)}}{\sqrt{V(z_2)}}$$~~

$$\text{Corr}(\hat{x}_2, \hat{z}_1)$$

$$= \text{Corr}(\hat{x}_2, \hat{x}_3 - \hat{x}_2)$$

$$= r_{23} - \rho_{12}$$

$$\text{Corr}(\hat{x}_1, \hat{z}_1)$$

$$= \text{Corr}(\hat{x}_1, \hat{x}_3)$$

$$= \rho_{13} - \rho_{12}$$

$$\text{Corr}(\hat{x}_1, \hat{z}_2)$$

$$= \rho_{14} - \rho_{12}$$



$$x_1 = a + b_1' x_2 + b_3 x_3 + b_4 x_4$$

$$a = \bar{x}_1 - b_1' \bar{x}_2 - b_3 \bar{x}_3 - b_4 \bar{x}_4$$

$$b_1' = 1 - b_3 - b_4$$

$$b_3 = -\frac{R_{1,3}}{R_{1,1}} \frac{\delta_1}{\delta_3} = -0.2189$$

$$b_4 = -\frac{R_{1,4}}{R_{1,1}} \frac{\delta_1}{\delta_4} = -0.0853$$

$$a \cdot b_1' = 1.3042$$

$$a = 56.806 - (1.3042 \times 68.513) - (0.2189 \times 75.48) - (0.0853 \times 55.23) = -11.3149$$

$$\text{Cov}(e_{1,34\dots p}, e_{2,34\dots p})$$

$$= \frac{1}{n} \sum_{\alpha=1}^m (e_{1,34\dots p\alpha} \cdot e_{2,34\dots p\alpha}) \quad \left[ \text{as } \bar{e}_{1,34\dots p} = 0 \text{ and } \bar{e}_{2,34\dots p} = 0 \right]$$

$$= \frac{1}{n} \sum_{\alpha=1}^m \left( x_{1\alpha} - \bar{x}_1 + \frac{R_{1,3}^{(2)}}{R_{1,1}^{(2)}} \frac{\delta_1}{\delta_3} (x_{3\alpha} - \bar{x}_3) \right)$$

$$+ \left( \dots + \frac{R_{1,p}^{(2)}}{R_{1,1}^{(2)}} \frac{\delta_1}{\delta_p} (x_{p\alpha} - \bar{x}_p) \right) e_{2,34\dots p\alpha}$$

$$= \frac{1}{n} \sum (x_{1\alpha} - \bar{x}_1) e_{2,34\dots p\alpha}$$

$$\text{As } \sum (x_{i\alpha} - \bar{x}_i) e_{2,34\dots p\alpha} = 0$$

for  $i = 3, 4, \dots, p$

$\text{Cov}(x_i, y)$

$$= \frac{1}{n} \sum_{i=1}^n x_i (y_i - \bar{y})$$

$$= \frac{1}{n} \sum (x_i - \bar{x})(y_i - \bar{y})$$

due to normal eq  $\frac{n}{n}$  related to

$$x_2 = a^* + b_3^* x_3 + \dots + b_p^* x_p$$

$$= \frac{1}{n} \sum (\chi_{1\alpha} - \bar{x}_1) \cdot e_{2,34 \dots p \alpha}$$

$$= \frac{1}{n} \sum_{\alpha=1}^n (\chi_{1\alpha} - \bar{x}_1) \left( \chi_{2\alpha} - \bar{x}_2 + \frac{R_{2,3}^{(1)}}{R_{2,2}^{(1)}} \frac{\delta_2}{\delta_3} \right.$$

$$\left. (\chi_{3\alpha} - \bar{x}_3) + \dots + \frac{R_{2,p}^{(1)}}{R_{2,2}^{(1)}} \frac{\delta_2}{\delta_p} (\chi_{p\alpha} - \bar{x}_p) \right)$$

$$= \frac{1}{n} \sum_{\alpha=1}^n (\chi_{1\alpha} - \bar{x}_1) (\chi_{2\alpha} - \bar{x}_2)$$

$$+ \frac{R_{2,3}^{(1)}}{R_{2,2}^{(1)}} \frac{\delta_2}{\delta_3} \frac{1}{n} \sum_{\alpha} (\chi_{1\alpha} - \bar{x}_1) (\chi_{3\alpha} - \bar{x}_3)$$

$$+ \dots + \frac{R_{2,p}^{(1)}}{R_{2,2}^{(1)}} \frac{\delta_2}{\delta_p} \frac{1}{n} \sum_{\alpha} (\chi_{1\alpha} - \bar{x}_1) (\chi_{p\alpha} - \bar{x}_p)$$

$$= r_{12} \delta_1 \delta_2 + \frac{R_{2,3}^{(1)}}{R_{2,2}^{(1)}} \frac{\delta_2}{\delta_3} \cdot r_{13} \delta_1 \delta_3 + \dots + \frac{R_{2,p}^{(1)}}{R_{2,2}^{(1)}} \frac{\delta_2}{\delta_p}$$

$$= \frac{\delta_1 \delta_2}{R_{2,2}^{(1)}} \left[ r_{12} R_{2,2}^{(1)} + r_{13} R_{2,3}^{(1)} + \dots + r_{1p} R_{2,p}^{(1)} \right]$$

The expression inside the bracket is the determinant obtainable from  $R^{(1)}$  by replacing its first row  $(r_{12}, r_{13}, \dots, r_{1p})$  by  $(r_{12}, r_{13}, \dots, r_{1p})$

$$x_2 = a + b_3 x_3 + \dots + b_p x_p$$

$$\sum (\chi_{2\alpha} - \bar{x}_2 - b_3 (\chi_{3\alpha} - \bar{x}_3) - \dots - b_p (\chi_{p\alpha} - \bar{x}_p))$$

$$\frac{\partial L}{\partial b_3} = (-2) \sum_{\alpha} (\chi_{3\alpha} - \bar{x}_3) e_{2,3 \dots p \alpha}$$

$$\equiv \frac{\delta_1 \delta_2}{R_{2,2}^{(1)}} \begin{vmatrix} r_{12} & r_{13} & \dots & r_{1p} \\ r_{32} & r_{33} & \dots & r_{3p} \\ \vdots & \vdots & \ddots & \vdots \\ r_{p2} & r_{p3} & \dots & r_{pp} \end{vmatrix} \quad R_{2,2}^{(1)} = \begin{vmatrix} r_{11} & r_{12} & r_{13} & \dots & r_{1p} \\ r_{21} & r_{22} & r_{23} & \dots & r_{2p} \\ r_{31} & r_{32} & r_{33} & \dots & r_{3p} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ r_{p1} & r_{p2} & r_{p3} & \dots & r_{pp} \end{vmatrix}$$

$= -\frac{\delta_1 \delta_2}{R_{2,2}^{(1)}} R_{2,1}$  where  $R_{2,1}$  is the cofactor of  $r_{21}$  in  $R$

Finally

$$\frac{r_{12} r_{21} \dots r_{1p} r_{2p} \dots r_{p2} r_{pp}}{R_{2,2}^{(1)}} = \frac{-\frac{\delta_1 \delta_2}{R_{2,2}^{(1)}} R_{2,1}}{\sqrt{\frac{\delta_1^2 [R_{1,1}^{(2)}]}{R_{1,1}^{(2)}} \cdot \frac{\delta_2^2 [R_{2,2}^{(1)}]}{R_{2,2}^{(1)}}}}$$

$$= \frac{R_{2,1} [R_{1,1}^{(2)}]^{1/2}}{[R_{2,2}^{(1)}]^{1/2} [R^{(2)}]^{1/2} [R^{(1)}]^{1/2}}$$

But,  $R^{(1)} = R_{1,1}$

and  $R^{(2)} = R_{2,2}$

$$\frac{R_{2,1} (R_{1,1}^{(2)})^{1/2}}{\sqrt{R_{1,1} R_{2,2} (R_{2,2}^{(1)})^{1/2}}} = \frac{R_{2,1}}{\sqrt{R_{1,1} R_{2,2}}}$$

$$R = \begin{pmatrix} r_{11} & r_{12} & r_{13} & r_{14} & \dots & r_{1p} \\ r_{21} & r_{22} & r_{23} & r_{24} & \dots & r_{2p} \\ r_{31} & r_{32} & r_{33} & r_{34} & \dots & r_{3p} \\ \vdots & \vdots & \vdots & \vdots & \ddots & \vdots \\ r_{p1} & r_{p2} & r_{p3} & r_{p4} & \dots & r_{pp} \end{pmatrix}$$

Cofactor of  $R_{1,1}$  i.e.  $r_{11}$

$$\text{As } \begin{pmatrix} R_{1,1}^{(2)} \\ R_{1,1} \end{pmatrix} = \begin{pmatrix} (1) \\ R_{2,2} \end{pmatrix}$$

Similarly  
first find  
 $R^{(1)}$   
then cofactor  
of  $r_{22}$

Partial correlation coefficient lies within -1 to 1  
as the numerator (covariance) can be positive/  
negative unlike multiple correlation where the  
numerator (covariance  $(x_1, x_{1.23\dots p})$ ) turns  
be  $V(x_{1.23\dots p})$

Standard error of estimate (Residual variance)

$$s_{1.23\dots p}^2 = V(e_{1.23\dots p}) = \frac{|R^{(1)}|}{R_{1,1}^{(2)}} s_1^2 = \frac{R_{2,2}}{R_{1,1}^{(2)}}$$

$$s_{2.34\dots p}^2 = V(e_{2.34\dots p}) = \frac{|R^{(2)}|}{R_{2,2}^{(1)}} s_2^2 = \frac{R_{1,1}}{R_{2,2}^{(1)}}$$

$$\text{and } r_{12.34\dots p} = \frac{-R_{2,1}}{\sqrt{R_{1,1} R_{2,2}}}$$

$$b_{12.34\dots p} = - \frac{R_{1,2}}{R_{1,1}} \frac{\delta_1}{\delta_2}$$

Relationship between partial regression coeff, partial corr. coeff, residual variances

$$\frac{\delta_{1.34\dots p}^2}{\delta_{2.34\dots p}^2} = \frac{R_{2,2} \delta_1^2}{R_{1,1} \delta_2^2}$$

$$\Rightarrow \frac{\delta_{1.34\dots p}}{\delta_{2.34\dots p}} = \sqrt{\frac{R_{2,2}}{R_{1,1}}} \cdot \frac{\delta_1}{\delta_2}$$

$$\frac{\delta_{1.34\dots p}}{\delta_{2.34\dots p}} \frac{\sqrt{R_{1,1}}}{\sqrt{R_{2,2}}} = \frac{\delta_1}{\delta_2}$$

$$\therefore b_{12.34\dots p} = - \frac{R_{1,2}}{R_{1,1}} \cdot \frac{\delta_{1.34\dots p}}{\delta_{2.34\dots p}} \sqrt{\frac{R_{1,1}}{R_{2,2}}}$$

$$= - \frac{R_{1,2}}{\sqrt{R_{1,1} R_{2,2}}} \cdot \frac{\delta_{1.34\dots p}}{\delta_{2.34\dots p}}$$

$$= r_{12.34\dots p} \frac{\delta_{1.34\dots p}}{\delta_{2.34\dots p}}$$

Here  $R_{1,2} = R_{2,1}$   
as  $R$  is symmetric matrix

• Properties of residuals

$$x_i = x_{i.23\dots p} + e_{i.23\dots p}$$

$$= a + b_2 x_{i2} + \dots + b_p x_{ip}$$

$$= \bar{x}_i - b_2 \bar{x}_2 - \dots - b_p \bar{x}_p + b_2 x_{i2} + \dots + b_p x_{ip}$$

From 2nd to  
p-th normal  
equation

$$\sum_{\alpha=1}^n (x_{i\alpha} - \bar{x}_i) e_{i.23\dots p\alpha} = 0$$

$$\Rightarrow \sum x_{i\alpha} e_{i.23\dots p\alpha} = 0$$

From the partial regression concept

$$\textcircled{1} \begin{cases} x_i = x_{i.34\dots p} + e_{i.34\dots p} \\ = a^* + b_3^* x_{i3} + \dots + b_p^* x_{ip} + e_{i.34\dots p} \end{cases}$$

$$\textcircled{2} \begin{cases} x_2 = x_{2.34\dots p} + e_{2.34\dots p} \\ = a^{\circ} + b_3^{\circ} x_{23} + \dots + b_p^{\circ} x_{2p} + e_{2.34\dots p} \end{cases}$$

from ①  $\sum_{\alpha=1}^n (x_{i\alpha} - \bar{x}_i) e_{i.34\dots p\alpha} = 0, i=3,4,\dots,p$

from ②  $\sum_{\alpha=1}^n (x_{i\alpha} - \bar{x}_i) e_{2.34\dots p\alpha} = 0, i=3,4,\dots$

$$b_2 = b_{2.34\dots p}$$

$$\sum_{\alpha=1}^n e_{i.23\dots p\alpha}$$

$$= \sum_{\alpha=1}^n e_{i.23\dots p\alpha} + e_{i.23\dots p\alpha}$$

$$= \sum (x_{i\alpha} - \bar{x}_i - b_{2.34\dots p} (x_{i2} - \bar{x}_2) - \dots - b_{1p.23\dots p-1} (x_{ip} - \bar{x}_p)) e_{i.23\dots p\alpha}$$

$$e_{i.23\dots p\alpha}$$

$$L = \sum (x_{i\alpha} - a - b_2 x_{i2} - \dots - b_p x_{ip})^2$$

$$\frac{\partial L}{\partial b_2} = 0;$$

$$\Rightarrow \sum (x_{i\alpha} - a - b_2 x_{i2} - \dots - b_p x_{ip}) x_{i2} = 0$$

$$\Rightarrow \sum (x_{i\alpha} - a - b_2 x_{i2} - \dots - b_p x_{ip}) = 0$$

$$\Rightarrow \sum e_{i.23\dots p}$$

$$= \sum (\chi_{1\alpha} - \bar{\chi}_1) e_{1.23\dots p\alpha} - b_{12.31\dots p} \sum (\chi_{2\alpha} - \bar{\chi}_2) e_{1.23\dots p\alpha} - \dots - b_{1p.23\dots p-1} \sum (\chi_{p\alpha} - \bar{\chi}_p) e_{1.23\dots p\alpha}$$

$$= \sum_{\alpha} (\chi_{1\alpha} - \bar{\chi}_1) e_{1.23\dots p\alpha} \quad \left[ \begin{array}{l} \text{Due to normal equation} \\ \text{Other terms will be zero} \\ \text{from the multiple regression} \\ \text{equation} \end{array} \right]$$

$$= \sum_{\alpha=1}^n \chi_{1\alpha} \cdot e_{1.23\dots p\alpha}$$

$$= \sum_{\alpha} (\chi_{1\alpha} - a' - b'_{12.34\dots p-1} \chi_{2\alpha} - b'_{13.24\dots p-1} \chi_{3\alpha} - \dots - b'_{1p-1.23\dots p-2} \chi_{p-1\alpha}) e_{1.23\dots p\alpha}$$

least square  
Think on a multiple regression line of  $\chi_1$  on  $\chi_2, \chi_3, \dots, \chi_{p-1}$

$$\chi_1 = a' + b'_2 \chi_2 + b'_3 \chi_3 + \dots + b'_{p-1} \chi_{p-1}$$

$$\sum_{\alpha=1}^n e_{1.23\dots p-1\alpha} = 0$$

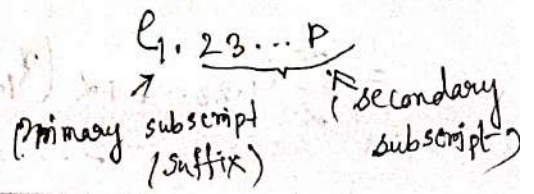
$$\sum_{\alpha=1}^n (\chi_{i\alpha} - \bar{\chi}_i) e_{1.23\dots p-1\alpha} = 0 \quad i = 2, 3, \dots, p-1$$

$$\rightarrow \equiv \sum_{\alpha=1}^n e_{1.23\dots p-1\alpha} \cdot e_{1.23\dots p\alpha}$$

$$= \sum e_{1.23\dots p-2\alpha} \cdot e_{1.23\dots p-1\alpha}$$

### Statement

The sum of product of any two residuals in which all the subscripts after dot of the first term occur among the subscripts of the second is unaltered if we omit any secondary subscript of the first.



Result  $p$ th order multiple correlation than  $(p-1)$ st order multiple correlation.

$$r_{1.23\dots p}^2 \geq r_{1.23\dots p-1}^2$$

In general

$$r_{1.23\dots p}^2 \geq r_{1.23\dots p-1}^2 \geq r_{1.23\dots p-2}^2 \geq \dots \geq r_{1.2}^2$$

Proof  $\Rightarrow V(\ell_{1.23\dots p}) = \delta_{1.23\dots p}^2 = \frac{1}{n} \sum_{\alpha=1}^n \ell_{1.23\dots p \alpha}^2$

$$\Rightarrow n \cdot \delta_{1.23\dots p}^2 = \sum_{\alpha=1}^n \ell_{1.23\dots p \alpha}^2$$

$$\begin{aligned} x_{1\alpha} &= a^0 + b_2^0 x_{2\alpha} + \dots + b_{p-1}^0 x_{p-1\alpha} \\ &+ \ell_{1.23\dots p-1 \alpha} \\ a &= \bar{x}_{1\alpha} - b_2 \bar{x}_2 - \dots - b_{p-1} \bar{x}_{p-1} \\ \sum (x_{i\alpha} - \bar{x}_i) \ell_{1.23\dots p-1 \alpha} &= 0 \end{aligned}$$

$$= \sum_{\alpha=1}^n \ell_{1.23\dots p-1 \alpha} \cdot \ell_{1.23\dots p \alpha}$$

$$= \sum_{\alpha=1}^n (x_{1\alpha} - \bar{x}_1) \ell_{1.23\dots p \alpha}$$

$$= \sum_{\alpha=1}^n (x_{1\alpha} - a' - b_2' x_{2\alpha} - \dots - b_{p-1}' x_{p-1\alpha}) \ell_{1.23\dots p \alpha}$$

$$= \sum_{\alpha=1}^n (x_{1\alpha} - \bar{x}_1 + b_2' \bar{x}_2 + \dots + b_{p-1}' \bar{x}_{p-1} - b_2' x_{2\alpha} - \dots - b_{p-1}' x_{p-1\alpha}) \ell_{1.23\dots p \alpha}$$

$$= \sum (x_{1\alpha} - \bar{x}_1) \ell_{1.23\dots p \alpha} - b_2' \sum (x_{2\alpha} - \bar{x}_2) \ell_{1.23\dots p \alpha}$$

$$- \dots - b_{p-1}' \sum (x_{p-1\alpha} - \bar{x}_{p-1}) \ell_{1.23\dots p \alpha} \rightarrow \text{zero}$$

Again  $n \cdot \delta_{1.23\dots p}^2 = \sum_{\alpha} \ell_{1.23\dots p-1 \alpha} \cdot \ell_{1.23\dots p \alpha}$

$$= \sum_{\alpha} \ell_{1.23\dots p-1 \alpha} (x_{1\alpha} - \bar{x}_1 - b_{12.3\dots p} (x_{2\alpha} - \bar{x}_2) - b_{13.2\dots p} (x_{3\alpha} - \bar{x}_3) - \dots - b_{1p.23\dots p-1} (x_{p\alpha} - \bar{x}_p))$$



$$= \sum_{\alpha} (x_{1\alpha} - \bar{x}_1) \epsilon_{1.23 \dots p-1 \alpha} - 0 - 0 \dots - 0 - b_{1p.23 \dots p-1} \sum (x_{1\alpha} - \bar{x}_1) \epsilon_{1.23 \dots p-1 \alpha}$$

$$= n \cdot \delta_{1.23 \dots p-1}^2 \quad [\text{From 13}]$$

$$- b_{1p.23 \dots p-1} \sum (x_{p\alpha} - \bar{x}_p) \epsilon_{1.23 \dots p-1 \alpha}$$

$$= n \cdot \delta_{1.23 \dots p-1}^2 - b_{p.23 \dots p-1} \sum (x_{p\alpha} - \bar{x}_p) \epsilon_{1.23 \dots p-1 \alpha} - b_{p2.34 \dots p-1} (x_{2\alpha} - \bar{x}_2) \epsilon_{1.23 \dots p-1 \alpha} - \dots - b_{p,p1.23 \dots p-2} (x_{p-1\alpha} - \bar{x}_{p-1}) \epsilon_{1.23 \dots p-1 \alpha}$$

$$x_p = a^0 + b_1^0 x_1 + b_2^0 x_2 + \dots + b_{p-1}^0 x_{p-1} + \epsilon_{p.12 \dots p-1}$$

$$a^0 = \bar{x}_p - b_1^0 \bar{x}_1 - \dots - b_{p-1}^0 \bar{x}_{p-1}$$

Rough

$$\sum (x_{i\alpha} - \bar{x}_i) \epsilon_{p.12 \dots p-1 \alpha}, \quad i = 1(1)p-1$$

miss  $\epsilon_{1.23 \dots p-1 \alpha}$  term into 1 and miss  $\epsilon_{p.23 \dots p-1 \alpha}$  into 2

$$= n \delta_{1.23 \dots p-1}^2 - b_{1p.23 \dots p-1} \sum \epsilon_{p.23 \dots p-1 \alpha} \epsilon_{1.23 \dots p-1 \alpha}$$

$$= n \cdot \delta_{1.23 \dots p-1}^2 - n b_{1p.23 \dots p-1} \sum \epsilon_{1p.23 \dots p-1} \delta_{p.23 \dots p-1} \delta_{1.23 \dots p-1}$$

(\*)

Rough

$$r_{12.34 \dots p}^2 = \frac{\text{cov}(\epsilon_{1.34 \dots p}, \epsilon_{2.34 \dots p})}{\delta_{1.34 \dots p} \cdot \delta_{2.34 \dots p}}$$

$$r_{12.3 \dots p}^2 \times \delta_{1.34 \dots p} \times \delta_{2.34 \dots p} = \frac{1}{n} \sum \epsilon_{1.34 \dots p \alpha} \cdot \epsilon_{2.34 \dots p \alpha}$$

$$\Rightarrow n r_{12.3 \dots p}^2 \times \delta_{1.34 \dots p} \times \delta_{2.34 \dots p} = \sum \epsilon_{1.34 \dots p \alpha} \cdot \epsilon_{2.34 \dots p \alpha}$$

$$\begin{aligned}
 & \textcircled{*} = \sigma^2 \delta_{1,2,3 \dots p-1}^2 - \sigma^2 r_{1p,2,3 \dots p-1}^2 \frac{\delta_{1,2,3 \dots p-1}^2}{\delta_{p,2,3 \dots p-1}^2} \cdot r_{1p,2,3 \dots p-1}^2 \\
 & \hspace{15em} \delta_{p,2,3 \dots p-1}^2 \delta_{1,2}
 \end{aligned}$$

$$\Rightarrow \sigma^2 \delta_{1,2,3 \dots p}^2 > \sigma^2 \delta_{1,2,3 \dots p-1}^2 (1 - r_{1p,2,3 \dots p-1}^2)$$

Since,  $0 < r_{1p,2,3 \dots p-1}^2 < 1$

$\therefore 0 < 1 - r_{1p,2,3 \dots p-1}^2 < 1$

$$\Rightarrow \delta_{1,2,3 \dots p}^2 \leq \delta_{1,2,3 \dots p-1}^2$$

[ Residual variance decreases as the number of variables under model increases ]

Practical

② Examine the following correlation matrix for internal consistency -

$$\begin{pmatrix}
 1 & -0.175 & 0.812 & 0.463 \\
 0.175 & 1 & 0.712 & 0.139 \\
 0.812 & 0.712 & 1 & 0.268 \\
 0.463 & 0.139 & 0.268 & 1
 \end{pmatrix}$$

Ans

And  $r_{1.234}^2$

③ Suppose a computer has found for a given set of values  $x_1, x_2, x_3$ ;

$$r_{12} = 0.91, \quad r_{13} = 0.33, \quad r_{23} = 0.81$$

whether  
Examine the computer is free from error.

Method (i)  $|R|$  is positive

(ii)  $r_{1.23}$  is positive.

(If negative then not free from error)

Answer

② Given correlation matrix,

$$R = \begin{pmatrix} 1 & 0.175 & 0.812 & 0.463 \\ 0.175 & 1 & 0.712 & 0.139 \\ 0.812 & 0.712 & 1 & 0.268 \\ 0.463 & 0.139 & 0.268 & 1 \end{pmatrix}$$

$\therefore$  Now, we have to find  $R^2$  i.e.  $r_{1.234}^2$ , that is coefficient of determination.

$$\text{where } r_{1.234} = \left( 1 - \frac{|R|}{R_{1,1}} \right)^{1/2}$$

$$|R| = -0.01530499$$

$$R_{1,1} = 0.4549578$$

$$\therefore r_{1.234} = \left( 1 + \frac{0.01530499}{0.4549578} \right)^{1/2}$$

$$\therefore r_{1.234}^2 = 1.08263$$

we know,  $0 < r \leq 1$  but here  $r_{1,2,3}$   
So, there is no internal consistency.

(3) Given correlation matrix

$$R = \begin{pmatrix} 1 & 0.91 & 0.33 \\ 0.91 & 1 & 0.81 \\ 0.33 & 0.81 & 1 \end{pmatrix}$$

$$r_{12} = 0.91$$

$$r_{13} = 0.33$$

$$r_{23} = 0.81$$

$$\text{Now, } |R| = -0.106614$$

$$r_{1,2,3} = \left( 1 - \frac{|R|}{R_{1,1}} \right)^{1/2}$$

$$R_{1,1} = 0.3439$$

$$= 1.14455867$$

We see the  $|R|$  is negative and also multiple correlation coefficient  $r_{1,2,3}$  is 1.14455867 i.e. it does not belong the range 0 to 1 not  
So, we may conclude that the computer is free from error.